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Regular rates of popular culture change reflect random copying

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Abstract

Almost by definition, "popular culture" reflects the effects of most people imitating those around them. At the same time, trends and fashions are constantly changing, with future outcomes potentially irrational and nearly impossible to predict. A simple null model, which captures these seemingly conflicting tendencies of conformity and change, involves the random copying of cultural variants between individuals, with occasional innovation. Here, we show that the random-copying model predicts a continual flux of initially obscure new ideas (analogous to mutations) becoming highly popular by chance alone, such that the turnover rate on a list of most popular variants depends on the list size and the amount of innovation but not on population size. We also present evidence for remarkably regular turnover on "pop charts"—including the most popular music, first names, and dog breeds in 20th-century United States—which fits this expectation. By predicting parametric effects on the turnover of popular fashion, the random-copying model provides an additional means of characterizing collective copying behavior in culture evolution.

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Keywords: Neutral theory; Random genetic drift; Pop music; Markets; Cultural evolution; Baby names; Cultural transmission; Power law; Fashion; Purebred dogs; Random copying

1. Introduction

As Boyd and Richerson (1985, p. 33) defined over 20 years ago, "culture is information capable of affecting individuals' phenotypes which they acquire from other conspecifics by teaching or imitation." Imitation is arguably the simplest form of culture transmission, termed unbiased transmission by Boyd and Richerson, which occurs when each individual acquires his or her behavior simply by copying from another individual within the population. Copying is a predominant human behavior (e.g., Gergely, Bekkering, & Király, 2004; Iacoboni et al., 1999) and is shared among primates (cf. Subiaul, Cantlon, Holloway, & Terrace, 2004). It can, thus, be useful to assume, as a null hypothesis in certain instances of social choice, that people

simply copy each other at random. In cases where choices 39 have intrinsic value with respect to one another, it makes 40 more sense to assume that cost-benefit decisions are made 41 independently, with conformity potentially among the biases 42 in making those decisions (e.g., Boyd & Richerson, 1985, 43 2005; Gintis, in press; Henrich, 2001, 2004; Henrich & 44 Boyd, 2001; McElreath, Boyd, & Richerson, 2003; 45 Shennan, 2002). This distinction is crucial to the nature of 46 collective human behavior, in anything from voting, to 47 corporate boardrooms, to deciding on a hunting strategy, as 48 copying can tend toward baseless decisions, whereas 49 independent decision making may lead to a rational, 50 collective "wisdom" of a group (Surowiecki, 2004) and/or 51 optimal solutions through a process analogous to natural 52 selection (e.g., Crow & Aoki, 1982a, 1982b, 1984; Henrich, 53 Q2 2004). While there is a fairly large body of literature on 54 group norms that arise as a consequence of identifiable costs 55 and benefits of cultural traits, quantitative models of random 56 copying of neutral cultural traits are relatively less well 57 developed (see reviews by Eerkens & Lipo, 2005; Mesoudi, 58

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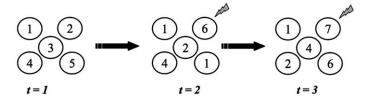


Fig. 1. A simple representation of the neutral-trait model. Shown are five individuals for three successive time steps. At each time step, we refresh the population with new individuals, and each is given a new copy of a variant (represented by numbers inside the circles). Each variant is assigned a new value by either (a) copying a randomly selected individual from the previous time step, with equal probability of choosing any individual, or (b) inventing a new variant (gray lightning bolts) with probability μ , the fraction of innovators among the N individuals.

Whiten, & Laland, in press). Here, we focus on a particular prediction of the random-copying model, not to explain all human behavior, but to help identify it when it arises and further characterize the consequences of copying in collective behavior.

As we have shown in previous studies (Hahn & Bentley, 2003; Herzog, Bentley, & Hahn, 2004), a highly useful null hypothesis for popular culture change can be a process of random copying between individuals, akin to the process of random genetic drift in population genetics. With its great potential for future modification and development, there are many ways in which the random-copying model, with the resources of population genetics theory to support and develop it, can make substantial contributions to social science. Large-scale shifts in popular preferences (e.g., fashions) offer insight into general mechanisms of cultural change (Lieberson, 2000). Whereas the collective effect of independent decisions may be a sensible equilibrium, random copying is unpredictable, with no tendency toward an optimum. For example, a recent Internet-based sociological experiment (Salganik, Dodds, & Watts, 2006) demonstrated that popular success in music markets is as much a matter of social influence as of quality. A model that has proven surprisingly robust in explaining shifts in tastes assumes simply that the majority of individuals randomly copy the choices of others, with occasional innovation (Bentley, Hahn, & Shennan, 2004). In population genetics, a formal model of random copying between generations with mutation is called the neutral model (Kimura & Crow, 1964). While developed to explain genetic variability, the neutral model has been effectively applied to ecological and cultural phenomena (e.g., Cavalli-Sforza & Feldman, 1981; Dunnell, 1978; Hubbell, 2001; Lang & Barlow, 1997; Lipo, Madsen, Dunnell, & Hunt, 1997; Neiman, 1995). It predicts that, inevitably, some variants will become highly popular simply due to imitation, not because they are in some way "better" than other variants. We have found that the assumption of random copying provides realistic predictions of the frequency distribution and change in frequency over time of such diverse phenomena as Neolithic pottery decorations (Bentley & Shennan, 2003), baby names (Hahn & Bentley, 2003), and dog breeds (Herzog et al., 2004).

The random-copying model assumes that there are N102 individuals, each characterized by a behavioral/stylistic 103 variant (Fig. 1). At each time step, we refresh the population

with N new individuals, and each is assigned a new variant by 115 either (a) copying a randomly selected individual from the 116 previous time step, with equal probability of choosing any 117 individual, or (b) inventing a new variant with probability μ . 118 In each time step, most of the N new individuals are copiers, 119 while a fraction μ are innovators (with μ being a dimension- 120 less fraction, not a rate per time—by analogy, if a regular 121 delivery of N oranges has 5% rotten oranges each week, the 122 5% is a fraction, not a rate). The joint product of these two 123 parameters, $N\mu$, provides a population-level measure of 124 variation. Using this parameter and other results, the neutral 125 model provides testable predictions concerning the change 126 over generations in the number and relative frequencies of 127 different variants (Gillespie, 1998).

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Computer simulations of the neutral model show that the 129 distribution of variant popularity levels (frequencies) fol- 130 lows a power law function for small values of the innovation 131 fraction μ , and we have found that this prediction provides a 132 fit to the distributions of modern cultural variant frequencies 133 remarkably well (Bentley et al., 2004; Hahn & Bentley, 134 2003), which fits the analytical predictions of Ewens (1972). 135 An additional prediction of the neutral model is that if we 136 follow a set of variants introduced in the same generation, 137 the average of their frequencies stays the same over time, 138 but the disparity (variance) in their frequencies increases 139 (Hahn & Bentley, 2003). This provides a quantitative 140 expectation that was used in a case study of registered 141 purebred dog breeds in the United States (Herzog et al., 142 2004) to identify Dalmatians as an exceptional case that 143 cannot be explained by simple random copying and, thus, to 144 attribute the sudden popularity increase of Dalmatians to the 145 rerelease of the Disney movie 101 Dalmatians.

Another implication of the random-copying model is the 147 consistency of change of variants or fashions. Here, we 148 show that the random-copying model also predicts a 149 regularity of turnover among particularly popular variants 150 (fashions). Modern cultural data are commonly available in 151 the form of "Top y" lists of popularity, which represent the 152 Top y highest-frequency variants. Several variables could 153 affect differential turnover rates, including the length of the 154 list (e.g., Top 10 vs. Top 40 songs), the rate at which new 155 variants appear, and the population size. Our goal was to 156 explore how the turnover rate on a Top y list, which we refer 157 to as z_y , is affected by the length y of the list, innovation 158 fraction μ , and the population size N.

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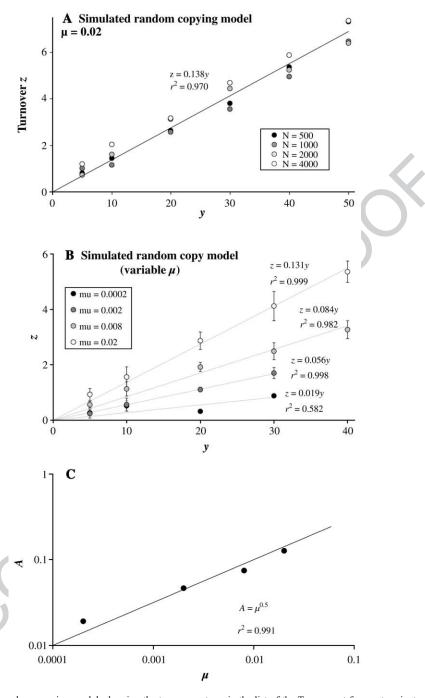


Fig. 2. Computer simulation of the random-copying model, showing the turnover rate z_y in the list of the Top y most-frequent variants, as a function of y. (A) Results for different numbers of individuals N, with the innovation fraction μ constant at 0.02 for all runs. (B) Results, for each value of μ , averaged from runs with N=250,500,1000, and 2000 (error bars showing $\pm 1\sigma$). (C) The slope, A, of each of the correlations in Panel B, versus μ .

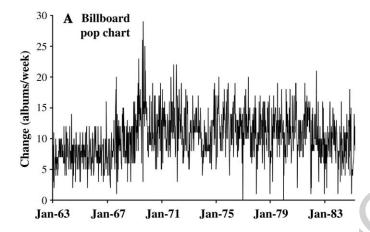
We are not aware of any direct analytical solution to this 161 problem since the sample includes only the most-frequent variants, which means that we cannot simply assume, as for 163 an entire population at equilibrium, that the innovation rate 164 balances the loss rate. Our approach, therefore, was to use 165 computer simulation (Bentley et al., 2004; Hahn & Bentley, 166 2003), by which we run the random-copying model 167 using different numbers of individuals, N, and innovation 168 fractions, μ . We then compared our simulation results to

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real-world data sets involving pop music, baby names, and 216 dog breeds in the 20th-century United States. 217

2. Methods 218

As described in detail previously (Bentley et al., 2004; 219 Hahn & Bentley, 2003), we used a simple computer simu- 220 lation of the neutral model written in a Java-based simulation 221 package called RePast (v 2.0, http://repast.sourceforge.net/). 222 R.A. Bentley et al. / Evolution and Human Behavior xx (2006) xxx-xxx



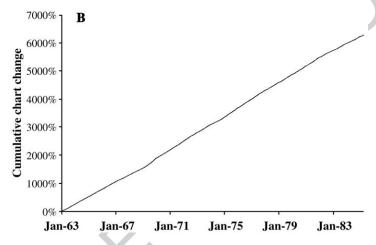


Fig. 3. (A) Weekly turnover on the Billboard Pop Chart, 1963–1985, in terms of the numbers of albums exiting the chart each week. (B) Cumulative change on the Pop Chart, in terms of the fraction of albums on the chart (z_y/y) exiting each week. The denominator y in calculating this fraction was variable, as the chart was expanded from 150 to 200 in mid-1967, and the actual value of y varies slightly from week to week (due to albums' shared positions on the chart, etc.). The turnover rate averaged 5.6% per week for over 20 years. Adapted from Bentley and Maschner (1999).

223 As represented schematically in Fig. 1, the simulation begins 224 with Í individuals that are assigned Í different variants, which are then subject to repeated copying and innovation (cf. 226 mutation). The simulation records the occurrence of every 227variant to appear in the population throughout the run. At 228 every time step, the I individuals are replaced with I new 229 individuals, the majority of which receives a variant copied at random from the previous time step, while the remaining 230 minority ($N\mu$, where μ is the innovation fraction) invents a novel variant ("innovation"). After running the simulation for 250 time steps to reach a quasi-equilibrium state, we recorded all the variants present in the population and their frequencies, for every other time step until Time Step 300 236 (25 total samples). At each sampling point, we recorded all the variants present and their frequencies. We then created Top y charts of varying sizes (y=5, 10, 20, 30, 40) for each of these samples. To determine the turnover rate among the Top y most-frequent variants, we ranked the variants by abundance for each sampled interval and then tabulated the number of new variants to enter the Top y chart relative to the 243 previously sampled interval.

279 3. Results

The simulations all show that, after the transient phase 280 of the first 250 time steps, the turnover rate z_y of a Top y list 281 (the number of new variants to enter the Top y chart relative to 282the previously sampled interval) finds a steady state, where it 283 fluctuates around a nominal average. In this steady state, as 284 Fig. 2A shows, the average turnover is linearly proportional 285 to the size of the list y (r^2 =.970). There is, therefore, a strong dependence of turnover rate on the length y of the Top y list.

Our simulation results also showed that under the 288 assumption of random copying, the turnover rate is 289 predicted to be independent of the population size N, as 290 varying N from 250 to 2000 in the simulations (Fig. 2A) had 291 no significant effect on the turnover rate (r^2 =.003). In sum, 292 the average turnover in the simulated random-copying 293 model can be described simply as

$$z_{v} = Ay, \tag{1}$$

where A is a constant, such that z_v is largely independent of 296 N. At the same time, the simulation results show that z_v 297

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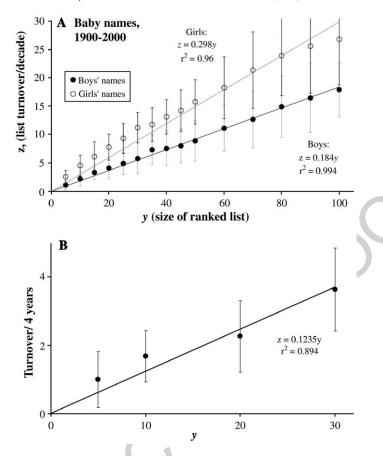


Fig. 4. Turnover rate of Top y charts, plotted against y, the size of the list, for (A) boys' names (filled circles) and girl's names (open circles) and for (B) dog breeds. For baby names, the turnover rates are per decade and averaged over the 20th century. For dog breeds, the turnover rates were calculated based on 4vear intervals. Error bars show $+1\sigma$.

298 does depend on the innovation fraction μ (Fig. 2B). In fact, as Fig. 2C shows, the constant A is simply proportional to μ , with an r^2 of .991 (p < .005). Hence, we find the simple relation:

$$z_{v} = y\sqrt{\mu}. (2)$$

An example that appears to exhibit characteristics of the 303 305 random-copying model involves the Billboard "Top 200" 306 Pop Chart (Bentley & Maschner, 1999; Whitburn, 1985)— 307 hereafter, the "Pop Chart"—a mainstream record of popular music in the United States. The number of new albums (or, equivalently, the number of exiting albums) per week on the Pop Chart jumped erratically about a nominal average— 311 rather like the "El Farol" problem of how many people visit a bar from night to night (Arthur, 1999)—of around 7– 313 8 albums per week from 1963 to 1967 when the Pop Chart 314 had 150 albums. This average then increased to around 11 315 albums per week after 1967, when the Pop Chart was 316 expanded to 200 albums (Fig. 3A). The change in the chart 317 size brought about a proportional change in the turnover 318 rate: when the Pop Chart was made 33% larger (from 150 to 200), the turnover rate increased by 38% (from about 8 to 319320 about 11 albums per week).

If we assume that a consistent fraction of the population 322 are innovators, then we can apply Eq. (1), which predicts

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that the fraction of Pop Chart turnover z_y/y , in terms of the 323 percentage change in the composition of the Pop Chart per 324 week, will be constant. Indeed, the Pop Chart turnover was 325 remarkably steady at 5.6% per week for over 20 years 326 (Fig. 3B). Unfortunately, we cannot systematically test 327 different values of y because our source (Whitburn, 1985) 328 only gives the date each album entered the Pop Chart and 329 the date it exited, rather than the position of each album 330 week to week. However, as Fig. 3B shows, there is no 331 visible change in the fractional turnover rate, z_y/y , when the 332 Pop Chart was expanded from 150 to 200 in mid-1967. Given that the population of the United States increased 334 almost 70% during this time, from about 150 million 335 in 1960 to about 250 million in 1990, and assuming a 336 commensurate increase in albums sold, it appears that the 337 steady turnover on the Pop Chart was independent of N, in 338 line with the random-copying model.

Our next real-world example involves popular baby 340 names, as recorded by the U.S. Social Security Administration (http://www.ssa.gov/OACT/babynames/) by ranking 342 the 1000 most common boys' and girls' names in each 343 decade of the 20th century. The rates at which new names 344 appeared on the list averaged 182±52 female names and 345 133±26 male names per decade. Taking these turnover 346 values as measures of z_{ν} , we find the turnover rate for 347

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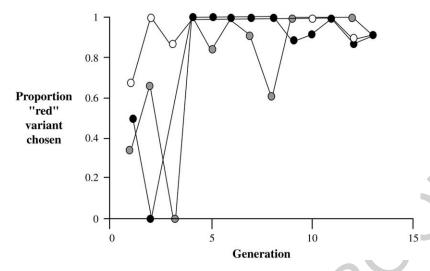


Fig. 5. Representative results of a cultural microevolution experiment by Baum et al. (2004), in which participants, in groups of four, solved puzzles that were coded either red or blue. In each generation, the player in the group who took the longest time in solving the puzzle was replaced by a new person. In the run shown in this figure, the payoffs for solving red puzzles was greater than those for blue, and as a result, red clearly becomes the predominant choice, although there is constant renewal of the population. The experiment shows how an intergenerational "tradition" results when there is a payoff advantage for a particular choice. Adapted from the study of Baum et al. (their Fig. 1C), showing the results of three typical sessions out of the six performed under the same parameters.

348 female names to be 1.37 times (182/133) that for male 349 names. The difference appears to be real and not by 350 statistical chance, as the rate for female names was higher 351 for every decade of the 20th century (Hahn & Bentley, 2003). The higher turnover rate for female names in each 353 decade implies more innovation in naming girls, as is clear 354 in other studies (see Fryer & Levitt, 2004; Levitt & Dubner, 355 2005, pp. 179–204; Lieberson, 2000).

As seen by the numbers of new names to enter the Top 20, 357 Top 100, Top 500, and Top 1000 lists in each decade, 358 the turnover rate clearly increases as y increases, for both boys' and girls' names. Averaged over all decades of the century, the resulting linear relationship with y (Fig. 4A) is as predicted by Eq. (1), for both girls' $(r^2=.96)$ and boys' $362 ext{ (}r^2 = .99 ext{)}$ names. According to Eq. (2), each slope (0.259 for 363 girls' names, 0.176 for boys' names) corresponds to the 364 square root of the innovation rate μ , yielding a century-365 averaged innovation parameter of 0.067 for girls' names and 366 0.031 for boys' names, for decadal sampling (since real-367 world events occur in time rather than orderly "generations," 368 the calculated innovation parameter is a relative measure 369 of the frequency of innovation for the sampling interval, 370 e.g., per decade).

Finally, we obtained data from the American Kennel 372 Club on the annual number of new puppy registrations in 373 the United States for all recognized breeds (Herzog et al., 374 2004). This is a large (a total of 52,806,268 registrations 375 from 1926 to 2004) and highly accurate index of the relative popularity of purebred dog breeds over the past five 377 decades. Using these data, we created multiple Top y lists 378 for each year since 1926 by listing and comparing the set of 379 top registered dog breeds in order of decreasing frequency 380 for the year. For each year in the study, we used the total 381 number of dogs registered as our measure of N. By

sampling lists of different sizes, y, at regular time intervals, 382 we then determined the turnover rate in each list per 4 years 383 (sampling every year would leave too many zeros in the 384 time series).

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The number of different dog breeds registered in the United States has increased from 73 breeds in 1926 to 150 387 breeds in 2004. Nonetheless, throughout this time period, 388 the pattern is as predicted by the random-copying model, 389 which is an increase in turnover rate, z_y , as y increases. As 390 Fig. 4B shows, the turnover rate z_v for dog breeds shows a 391 convincing $(r^2 = .894)$ linear relationship with y, and the 392 slope of 0.124 corresponds via Eq. (2) to an innovation 393 parameter μ of 0.015 for 4-year sampling.

4. Discussion 395

In accord with the evidence of copying behavior in 396 downloading music (Salganik et al., 2006), the random- 397 copying model provides a simple, parsimonious explanation 398 for the steady turnover of modern baby names, dog breeds, 399 and pop music albums over much of the 20th century. As is 400 often noted in the social sciences, many models can fit the 401 data, and we do not rule out the possibility that other models 402 could be devised to fit the patterns we have shown. We 403 advocate random copying as a null model firstly because it 404 appears to be the absolute simplest model capable of 405 replicating the data patterns at the societal scale and, 406 secondly, because its two mechanisms, innovation and 407 random copying, are the two most basic elements of 408 unbiased culture transmission as defined by Boyd and 409 Richerson (1985). In fact, there is growing evidence that in 410 situations where cultural transmission occurs predominantly 411 from one individual to another, a neutral or "random- 412 copying" model is the best null model against which 413

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414 real-world cultural variants can be compared (e.g., Bentley 415 et al., 2004; Bentley & Shennan, 2003, 2005; Hahn & 416 Bentley, 2003; Herzog et al., 2004; Lipo et al., 1997; Lynch, 1996; Neiman, 1995; Simkin & Roychowdhury, 2003). Use 418 of this model in future studies will make it easier to test for 419 more detailed effects, including those of race (e.g., Fryer & 420 Levitt, 2004), geography, and/or class. A recent study within 421 the publishing industry (Lulu.com, 2006), for example, shows a decreasing life expectancy of books on bestseller 423 lists since the 1950s (equivalent to an increasing turnover 424 rate), which, by comparison with the random-copying 425 model, would suggest an increase in innovation, perhaps 426 as books can be published more and more quickly in response to public topics and tastes.

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As Baum, Richerson, Efferson, and Paciotti (2004, p. 429 306) point out, while there is a wealth of research on individual-level mechanisms of social learning, there is a need for more discussion by social scientists of how cultural 432 traditions change over time at the population level. Applied 433 here on the societal scale, the random-copying model simply 434 assumes a constant proportion of innovators in the 435 population, as often assumed in epidemiological models applied to binary-choice culture change (e.g., Dodds & **Q4** 437 Watts, 2005; Watts, 2002). Of course, innovation, that is, creativity, is an enormous topic (e.g., Martindale, 1975, 1986, 1999) that we are not attempting to explain at the individual, psychological level. As we move outward in 440 scale, however, new quantitative effects emerge (Anderson, 442 1972), particularly with regard to the collective effects of 443 society (e.g., Ball, 2004; Barabási, 2005; Le Bon, 1896). Baum et al., for example, conducted an empirical investigation of "cultural microevolution" on the scales as small as 446 four individuals, which is an important bridge to the societal scale we are investigating here (we discuss their results below). At the societal scale, creative innovation generates what is effectively "quasi-random" variation as Martindale 450 (1986) described it.

An alternative to the random-copying model is clearly 452 some form of selection or that people choose the variant with the highest payoff with respect to some benefit (e.g., Baum et al., 2004; Boyd & Richerson, 1985; Gintis, in press; Henrich, 2001, 2006; Henrich et al., 2006). In the case of the phenomena we have investigated here—baby names, pop music, and dog breeds—we see no evidence for inherent benefits of one variant over another. As Baum et al. (2004) recently showed through experiments on groups of 460 four participants, the stronger the payoff is for choosing one particular variant over another, the stronger is the "tradition" that evolves in bringing that choice to be the most popular, which is passed on to the newcomers of each generation 464 (Fig. 5). Henrich (2001) and many others (e.g., Boyd & 465 Richerson, 1985; Dunnell, 1978; Gintis, in press) have made similar predictions—when variants are not neutral—then, we expect the most beneficial choice to rise to popularity and remain until a superior choice becomes available. Hence, 469 we would not expect constant, population-independent turnover among cultural variants if they were being selected 470 according to intrinsic value, although a future study of Top y 471 charts of nonneutral cultural variants could reveal intriguing 472 and unexpected aspects of their turnover.

Finally, it might be argued that none of the variants we 474 have studied is truly neutral for various reasons—certain 475 famous pop artists have clear advantages over newcomers, 476 for example, or people tire of old fashions in favor of new 477 ones (cf. pronovelty bias in Boyd & Richerson, 1985). This 478 is true, and it is the reason Herzog et al. (2004) could 479 identify Dalmatians as being selected among dog breeds 480 against a background of neutral evolution. To say that things 481 evolve neutrally means that, of the variants observed, all 482 behave in a neutral fashion. If they were all positively 483 selected compared to some unseen variant, they would still 484 all behave neutrally with respect to one another because 485 fitnesses are always relative. Hence, the inevitable fact that 486 newcomers have a lower fitness is not inconsistent with the 487 model. It would be interesting in the future to investigate the 488 effects of people getting tired of old fashions by intro- 489 ducing some sort of pronovelty bias among individuals 490 (cf. Shennan & Wilkinson, 2001) or by adding intrinsic 491 value to variants to introduce elements of selection, which 492 could even decay with time (cf. Dorogovtsev & Menedes, 493 2000). At this stage, however, the aim of the proponents of 494 the neutral model (Bentley et al., 2004; Bentley & Shennan, 495 2005; Hahn & Bentley, 2003; Herzog et al., 2004; Hubbell, 496 2001; Lipo et al., 1997; Neiman, 1995; Shennan & 497 Wilkinson, 2001) is still to establish the random-copying 498 model as an appropriate basis for making these added 499 alterations. For all its simplicity, the neutral model replicates 500 a remarkable range of patterns of cultural transmission. In 501 this study, our main aim is to show that while added rules 502 can always be imposed to engineer the results, the random- 503 copying model produces constant turnover on its own, 504 which we find unexpected, somewhat counterintuitive, and, 505 therefore, significant.

In conclusion, our simulations of the random-copying 507 model indicate that the population size N should not 508 significantly affect the turnover rate on the pop charts. 509 The time-averaged turnover rate in Top y charts is linearly 510 proportional to the chart size y and the square root of the 511 innovation fraction μ . Hence, while prediction of the next 512 big popular success may be impossible (Salganik et al., 513 2006), predicting the frequency distribution (Bentley et al., 514 2004; Hahn & Bentley, 2003) and turnover rate is relatively 515 straightforward. Since such regular turnover is not neces- 516 sarily expected when independent, rational decisions are 517 made, random copying may be identifiable in research of 518 markets and cultural change. The neutral model could also 519 be useful for assessing situations where copying and 520 continual, yet directionless, turnover may be undesirable, 521 as in politics or academic publishing. For these reasons 522 and more, further research on random copying should be 523 of high priority for the study of culture evolution and 524 collective behavior.

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